Population declines of tuna and relatives depend on their speed of life

M. J. Juan-Jordá1,2, I. Mosqueira3, J. Freire4 and N. K. Dulvy2

1AZTI Tecnalia, Herrera Kaia, Portualdea z/g, Pasai, Gipuzkoa 20110, Spain
2Earth to Ocean Research Group, Department of Biological Sciences, Simon Fraser University, Burnaby, British Columbia, Canada V5A 1S6
3European Commission, Joint Research Center, Institute for the Protection and Security of the Citizen, Maritime Affairs Unit G03, Ispra 21027, Italy
4Teamlabs, Calle de la Colegiata 9, Madrid 28012, Spain

Larger-bodied species in a wide range of taxonomic groups including mammals, fishes and birds tend to decline more steeply and are at greater risk of extinction. Yet, the diversity in life histories is governed not only by body size, but also by time-related traits. A key question is whether this size-dependency of vulnerability also holds, not just locally, but globally across a wider range of environments. We test the relative importance of size- and time-related life-history traits and fishing mortality in determining population declines and current exploitation status in tunas and their relatives.

We use high-quality datasets of half a century of population trajectories combined with population-level fishing mortalities and life-history traits. Time-related traits (e.g. growth rate), rather than size-related traits (e.g. maximum size), better explain the extent and rate of declines and current exploitation status across tuna assemblages, after controlling for fishing mortality. Consequently, there is strong geographical patterning in population declines, such that populations with slower life histories (found at higher cooler latitudes) have declined most and more steeply and have a higher probability of being overfished than populations with faster life histories (found at tropical latitudes). Hence, the strong, temperature-driven, latitudinal gradients in life-history traits may underlie the global patterning of population declines, fisheries collapses and local extinctions.

1. Introduction

Species are declining in abundance faster than ever in the history of the Earth [1,2]. Comparative studies of vulnerability to decline have contributed greatly to elucidate the underlying processes and patterns of species declines and extinction risk in a wide range of taxonomic groups [3–5]. The loss of biodiversity is not random [6,7]. Over the past decade, species vulnerability to declines and extinctions have been linked to two major factors: their exposure to a threatening process, and their intrinsic sensitivity based on their life-history traits, habitat preferences and behavioural ecology [3,8,9].

Marine fishes provide a unique opportunity to understand the intrinsic patterns and processes of decline and collapse. This is because exposure to the threatening process—fishing mortality—is routinely estimated in commercially important marine fish species. By contrast, in mammals, the variable population response of species is noted, but the local variation in hunting pressure among populations of a species is unknown and cannot be controlled for [3,10]. Once fishing mortality can be controlled for, it is apparent that life histories and demography relate to several measures of species vulnerability including declines, collapses, recoveries and threat status [9,11].

Specifically, maximum body size is a broadly reliable life-history correlate of population and species vulnerability to decline and extinction risk. Larger species tend to have declined more steeply and are under greater threat of extinction, than smaller-bodied species [12–15]. This is because larger-bodied fishes tend to have life-history strategies resulting in lower intrinsic rates of population increase and...
We calculated three metrics to describe population biomass trajectories, which we refer to as metrics of vulnerability to decline: (i) average annual rate of decline in adult biomass, (ii) total extent of decline in adult biomass, and (iii) current exploitation status defined using the $B_{\text{current}}/B_{\text{MSY}}$ ratio. $B_{\text{current}}/B_{\text{MSY}}$ is the ratio of the current adult biomass relative to the adult biomass that would provide the maximum sustainable yield (MSY) and determines whether a population is currently overfished ($B_{\text{current}} < B_{\text{MSY}}$) or not ($B_{\text{current}} > B_{\text{MSY}}$). For each population, we also calculated the average fishing mortality across all years ($F_{\text{average}}$) and divided it by the fishing mortality predicted to produce MSY ($F_{\text{MSY}}$), which we refer to as relative fishing mortality rate ($F_{\text{average}}/F_{\text{MSY}}$). We use this metric to control for the different fishing mortality rates that populations have experienced within their history of exploitation or at least as they started to be assessed (electronic supplementary material, S1). The relative fishing mortality rate metric was not available for three populations (North Pacific Albacore tuna, Northeast Pacific chub mackerel and Pacific bluefin tuna), and therefore we were not able to include these populations in our analyses (electronic supplement material, S1). Furthermore, we also excluded from the statistical analyses populations with increasing population trajectories (on average positive annual rates of change in adult biomass) as this study focuses on examining the interaction between life histories and fishing in determining population declines (electronic supplement material, S1).

We extracted the following life-history traits from a comprehensive life-history database [25,36]: maximum body size ($L_{\text{max}}$, cm), length and age-at-maturity ($L_{\text{m}}$, cm and $T_{\text{m}}$, years), longevity ($T_{\text{max}}$, years), growth rates described with the von Bertalanffy growth coefficient $k$ ($1/\text{year}$), the average absolute batch fecundity ($F_{\text{tot},\text{abs}}$), relative batch fecundity ($F_{\text{tot},\text{rel}}$), number of oocytes per gram and spawning interval ($SpW_{\text{max}}$, days) and duration ($SpW_{\text{max}}$, months). We report length-based estimates as fork lengths through our. For each life-history trait, we calculated a population-level estimate combining the life-history information from multiple studies carried out within their population distributions (electronic supplementary material, S2).

(b) Analyses
We fitted general and logistic linear models to 19 populations of scombrids to examine the relative importance of life histories and relative fishing mortality rates in determining their vulnerability to decline. These 19 populations have complete datasets including the three metrics of vulnerability to decline, a metric of relative fishing mortality and life-history traits (electronic supplementary material, S1 and S2). Specifically, we examined whether the small–large and slow–fast trait axes of life-history variation in scombrids explain their rate and extent of decline and current exploitation status, after controlling for the different relative fishing mortality rates experienced by populations. Maximum body size was used as a proxy to describe the small–large dimension of life-history variation and growth rate to describe the fast–slow dimension of life-history variation (electronic supplementary material, S2).

We assume that vulnerability to decline is a function of the exposure of a population to extrinsic threats, such as fishing mortality, coupled with the population’ intrinsic sensitivity based on their life histories. We tested the following hypotheses about intrinsic factors (life histories) while controlling for extrinsic factors (fishing)—declines will be more severe and exploitation status worse for: (H1) populations subject to higher fishing mortality rates; (H2) populations with larger body sizes after controlling for fishing mortalities; and (H3) populations with slower growth rates after controlling for fishing mortalities.

We fitted a general linear regression with normally distributed errors to model the two continuous dependent variables: the average annual rate of decline and total extent of decline in adult...
Table 1. Summary of explanatory generalized linear models evaluating the effects of relative fishing mortality rates and life history on three measures of vulnerability to decline: (a) rates of decline in adult biomass, (b) the extents of decline in adult biomass and (c) current exploitation status (probability of being overfished) in scombrid populations. (a,b) General linear regression with normally distributed errors and (c) logistic regression with binomial distributed errors. Models are sorted by AICc, which is the Akaike’s information criterion with a correction for small sample sizes; \(K\), number of parameters; \(I(\beta)\), the value of the maximized log-likelihood function; \(\Delta\text{AIC}_c = \text{AIC}_c - \text{min AIC}_c\), where \(\text{AIC}_c\) is the \(\text{AIC}_c\) for model \(i\), and \(\text{min AIC}_c\) is the smallest \(\text{AIC}_c\) value in the set of models; \(w_i\), the Akaike weights, expresses the relative likelihoods of candidate models, with the weight of any particular model varying from 0 (no support) to 1 (complete support) relative to the entire model set; \(R^2\), coefficient of determination.

<table>
<thead>
<tr>
<th>hypotheses</th>
<th>(K)</th>
<th>(I(\beta))</th>
<th>(\text{AIC}_c)</th>
<th>(\Delta\text{AIC}_c)</th>
<th>(w_i)</th>
<th>(R^2)</th>
</tr>
</thead>
<tbody>
<tr>
<td>(a) rate of decline in adult biomass</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>relative fishing mortality + growth rate</td>
<td>4</td>
<td>-13.48</td>
<td>37.82</td>
<td>0</td>
<td>0.57</td>
<td>0.31</td>
</tr>
<tr>
<td>relative fishing mortality</td>
<td>3</td>
<td>-15.76</td>
<td>39.11</td>
<td>1.29</td>
<td>0.3</td>
<td>0.12</td>
</tr>
<tr>
<td>historical fishing mortality + maximum size</td>
<td>4</td>
<td>-14.88</td>
<td>40.62</td>
<td>2.81</td>
<td>0.14</td>
<td>0.20</td>
</tr>
<tr>
<td>(b) extent of decline in adult biomass</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>relative fishing mortality + growth rate</td>
<td>4</td>
<td>-5.49</td>
<td>21.80</td>
<td>0.00</td>
<td>0.69</td>
<td>0.30</td>
</tr>
<tr>
<td>relative fishing mortality</td>
<td>3</td>
<td>-8.09</td>
<td>23.80</td>
<td>1.93</td>
<td>0.26</td>
<td>0.08</td>
</tr>
<tr>
<td>relative fishing mortality + maximum size</td>
<td>4</td>
<td>-8.04</td>
<td>26.90</td>
<td>5.08</td>
<td>0.05</td>
<td>0.08</td>
</tr>
<tr>
<td>(c) probability of being overfished</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>relative fishing mortality + growth rate</td>
<td>4</td>
<td>-8.88</td>
<td>25.00</td>
<td>0.00</td>
<td>0.75</td>
<td></td>
</tr>
<tr>
<td>relative fishing mortality</td>
<td>3</td>
<td>-11.56</td>
<td>27.70</td>
<td>2.69</td>
<td>0.20</td>
<td></td>
</tr>
<tr>
<td>relative fishing mortality + maximum size</td>
<td>4</td>
<td>-11.54</td>
<td>30.30</td>
<td>5.33</td>
<td>0.05</td>
<td></td>
</tr>
</tbody>
</table>

biomass. We fitted a logistic regression assigning a binomial error distribution and a logit link function to model the probability of the populations being overfished or not. We used an information-theoretic approach with Akaike’s information criterion corrected for small sample sizes (\(\text{AIC}_c\)) to evaluate all the candidate models and assign them relative strengths of evidence [37]. We determined the maximized log-likelihood for each candidate model \((i)\) and calculated the values for \(\text{AIC}_c\), \(\Delta\text{AIC}_c\) \((\Delta\text{AIC}_c = \text{AIC}_c - \text{min AIC}_c)\), where \(\text{AIC}_c\) is the \(\text{AIC}_c\) for model \(i\), and \(\text{min AIC}_c\) is the smallest \(\text{AIC}_c\) value in the set of models, and the Akaike weight \((w_i)\) [37]. We selected the best models (with largest Akaike weights) and calculated the standard error and the 95% confidence intervals (CIs) for each covariate to assess the effect size, and those variables whose CIs excluded zero were deemed to have a strong effect on the predictor variables [38]. We examined model diagnostics for heteroscedasticity, normality and independence of residuals [39].

Owing to issues of non-normality and non-constancy of variance (observed within the residual analyses), all the models were linearized by taking the natural logarithms of the variables. All data management, analyses and figures were done using the R statistical software v. 3.0.2 [40], including the R packages ‘MuMIn’ [41] and ‘ggplot2’ [42].

We conducted two sensitivity analyses to test the robustness of our choice of life-history proxies to describe the first two axes of life-history variation in scombrids, and found our analyses are robust to the choice of life-history proxies (electronic supplementary material, S2). Moreover, by examining the span of variation in the life-history traits and relative fishing mortalities against the three metrics of variability to decline, we also assessed whether populations with certain life histories have been preferentially targeted with high fishing intensities (electronic supplementary material, S3). We find scombrid populations, irrespective of their life histories, have been exposed to a wide range of relative fishing mortality rates. This exposure to wide-ranging fishing intensities and the diverse intrinsic sensitivities in scombrid populations allows testing for the combined effects of exposure to fishing and life histories in determining vulnerability to population declines.

3. Results

After controlling for the different relative fishing mortality rates experienced by each population, those populations with slower growth rates (rather than populations with larger body size) declined faster, to a greater extent, and more steeply, and are more likely to be currently overfished. We find there was a broad agreement between each of the three metrics of vulnerability to decline (rate, extent and exploitation status) and the life-history trait of growth rate, once the relative fishing mortality experienced by the populations was controlled for (table 1). We found greatest support for those models including growth rate and relative fishing mortality rate (both lowest \(\text{AIC}_c\) and highest weights \(w_i\) ranging from 0.57 to 0.75; table 1a–c). Thus, once fishing mortality was controlled for in the model, the growth rate of populations describing the slow–fast dimension of life histories better explained the variation in decline rate, extent and current exploitation status than maximum size representing the large–small life-history dimension. Moreover, we found weaker support \((w_i = 0.2–0.3)\) for the models that only included relative fishing mortality rates as an explanatory variable. Last, we found the weakest support \((w_i = 0.05–0.14)\) for the models including both maximum body size and relative fishing mortality rate as explanatory variables (table 1).

Those populations with slower growth rates, rather than populations with larger body size, are four times more likely to have experienced faster population declines (evidence ratio = 0.57/0.14), are 14 times more likely to have experienced larger extents of population declines (evidence ratio = 0.69/0.05) and are 15 times more likely to be currently overfished (evidence ratio = 0.75/0.05), after controlling for the different relative fishing mortality rates experienced by each population (table 1). Indeed, we find strong evidence for an effect of growth rate, as the 95% CIs of the effect sizes do not cross zero in all three measures of vulnerability (table 2). We also find evidence that,
Figure 1. Vulnerability to declines depends on somatic growth rates for scombrid populations after accounting for the exposure to the threatening process of fishing. (a) Predicted average annual rates of decline in adult biomass with separate lines set for two values of growth rates corresponding to the first (blue line, slow growth $k = 0.16$) and third (red line, fast growth $k = 0.34$) quartile values. (b) Predicted total extent of decline in adult biomass with separate lines set for two values of growth rates corresponding to the first (blue line) and third (red line) quartile values, $k = 0.16$ and $k = 0.34$, respectively. (c) Predicted probability of being overfished ($B_{current}/B_{MSY} < 1$) with separate lines set for two values of growth rates corresponding to the first (blue line, slow growth $k = 0.16$) and third (red line, fast growth $k = 0.34$) quartile values. Panels show regression lines and 95% CIs derived from (a,b) general linear models and (c) logistic regression models. Predictions with 95% confidence values correspond to the best models (the models with the largest Akaike weights, see table 1). Population codes are found in the electronic supplementary material, S1.
Table 2. Vulnerability to decline depends on somatic growth rates for scombrid populations after accounting for the relative fishing mortality rates populations have been exposed to. (Panels show selected best models (models with the largest Akaike weights from table 1) for each measure of vulnerability to decline. (a) The rates of decline in adult biomass, (b) the extent of decline in adult biomass, and (c) the current exploitation status (probability of being overfished) in scombrid populations. (a,b) General linear regression with normally distributed errors and (c) logistic regression with binomial distributed errors. The summary of the models includes the estimated coefficients, standard errors (s.e.) and the 95% CIs for each covariate.)

<table>
<thead>
<tr>
<th>parameter</th>
<th>estimate</th>
<th>s.e.</th>
<th>lower CI</th>
<th>upper CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>(a) rate of decline in adult biomass</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>relative fishing mortality</td>
<td>0.18</td>
<td>0.17</td>
<td>-0.16</td>
<td>0.52</td>
</tr>
<tr>
<td>growth rate</td>
<td>-0.51</td>
<td>0.24</td>
<td>-0.99</td>
<td>-0.03</td>
</tr>
<tr>
<td>(b) extent of decline in adult biomass</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>relative fishing mortality</td>
<td>0.07</td>
<td>0.11</td>
<td>-0.15</td>
<td>0.29</td>
</tr>
<tr>
<td>growth rate</td>
<td>-0.36</td>
<td>0.16</td>
<td>-0.68</td>
<td>-0.05</td>
</tr>
<tr>
<td>(c) probability of being overfished</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>relative fishing mortality</td>
<td>3.25</td>
<td>1.72</td>
<td>0.79</td>
<td>7.99</td>
</tr>
<tr>
<td>growth rate</td>
<td>-3.96</td>
<td>2.35</td>
<td>-10.36</td>
<td>-0.48</td>
</tr>
</tbody>
</table>

Parameter deemed significant as CI excludes 0.

4. Discussion

By focusing on a well-studied taxonomic group of species with detailed population trends and life-history data, we have shown that time-related traits describing the speed of life, rather than size-related traits, better explained the extent and rate of declines and current exploitation status of tunas and their relatives. At a first glance, our findings contrast strongly with the large majority of previous comparative studies revealing that large-bodied fish species tend to have declined more steeply and are under greater threat of extinction, than smaller-bodied species [12–14]. Instead, it is likely that our finding complements, rather than contradicts, these studies, and we explain how next.

There are a number of reasons why maximum body size might have been most frequently identified as the best life-history correlate of vulnerability in fishes. First, maximum size of the species tends to be most commonly available trait and often the only life-history trait available for testing, by comparison time-related traits have been less frequently tested [11,22]. Second, body-size-related measures are easier to gather accurately, with minimum observation error, whereas time-related traits must be estimated from numerous samples [36,43]. Third, some comparative studies were unable to control for (or only partially controlled for) the different fishing mortalities experienced by species [13,14,44]. This leads to a failure to adequately disentangle the relative importance of fishing and life histories with the result that the significance of maximum size may have been overestimated.

Aside from these quality and data availability issues, we find almost without exception that previous tests of life-history trait relationships with population trajectories while accounting for fishing mortality occurred within fish assemblages and were all conducted at local scales either in the European shelf seas [20,21,45] or the Pacific coral reefs [12,19,46]. Hence, within a single and relative homogeneous environment, maximum body size may be sufficient to rank the relative vulnerability of species to fishing exploitation within fish assemblages. Our study shows a geographical patterning in population declines with populations found at higher latitudes having declined most and more steeply than populations at lower latitudes after controlling for fishing mortality. This suggests time-related traits, such as growth rate, age at maturity and longevity, might be more suitable to rank population and species vulnerability to fishing exploitation across assemblages at larger geographical scales with larger environmental and temperature gradients. Yellowfin tuna is a good example. Although it is relatively large—up to 232 cm in length—it is a fast-growing and short-lived tropical species, and consequently it can cope with relatively high fishing mortality rates compared with the similar-sized temperate bluefin tunas [25]. We hypothesize that this is because life histories are locally adapted, most often to temperature and related environmental conditions [28,29,47].

To our knowledge, there are only four global scale tests of the link between life histories and vulnerability. First, a global
comparative study across freshwater and marine fishes showed body size best explained extinction risk in marine fishes [13,13]. Second, two global comparative studies across the taxonomic class Chondrichthyes—sharks, rays and chimaeras—similarly showed that body size, along with depth, best explained their extinction risk [14,48]. Like many other studies, these studies could not control for fishing mortality, nor were the authors able to test whether time-related traits were important, nor were traits geographically matched to population status. Third, a global analysis testing for life-history links to the proportion of populations (within each fish species) that have collapsed suggested life-history traits are not good correlates of species collapses [49]. Our advance, enabling us to reveal the importance of time-related traits, was to both control for fishing mortality as well as being able to geographically match life-history traits to population trends.

Our findings suggest it may be fruitful to better measure and account for external threats and their interaction with life-history traits, particularly time-related traits in comparative analysis of vulnerability. Time-related traits such as growth rates and age at maturity are increasingly identified in empirical studies as the primary correlates of the maximum population growth rates ($r_{max}$), a strong metric of species fitness and extinction risk [16,50,51]. Hence, time-related traits may serve as the ultimate correlate, whereas body size may serve as a useful proximate correlate of species vulnerability to decline, and extinction risk [16,18]. With increasing interest in the importance of scale in understanding patterns and processes of population vulnerability [10,52], future studies should also strive to bridge the gap between local and global processes of decline to better determine what aspects of species life-history histories predisposes them to be more susceptible to fishing pressure. The potential connection between biogeography, through temperature, and life histories may play an important role in determining the spatial patterning of fisheries yield, sustainability and vulnerability [30,53].

**Data accessibility.** The datasets supporting this article are available in the electronic supplementary material, S1 and S2.

**Authors’ contributions.** M.J.J.J., N.K.D., I.M. and J.F. designed and conducted the research and drafted the manuscript.

**Competing interests.** We have no competing interests.

**Funding.** M.J.J.J. was supported in part by a graduate scholarship from Caja Madrid Foundation, Spain, and an EU Marie Curie International Outgoing Fellowship—PIOF-GA-2013-628116. N.K.D. was supported by the Canada Research Chairs Program and Natural Science and Engineering Research Council Discovery and Accelerator Grant.

**Acknowledgements.** We gratefully acknowledge the constructive comments of Sebastian Pardo and Holly Kindsvater.

## References


